

Visualizing multivariate linear models in R

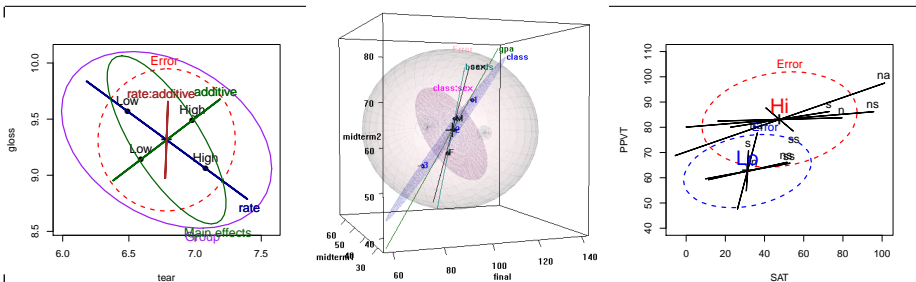
Michael Friendly¹ Matthew Sigal²

¹York University, Toronto

²Simon Fraser University

useR 2019

Toulouse, July 9–12, 2019



Outline

1 Background

- Overview
- Visual overview
- Data ellipses
- The Multivariate Linear Model

2 Hypothesis Error (HE) plots

- Motivating example
- Visualizing H and E variation
- MANOVA designs

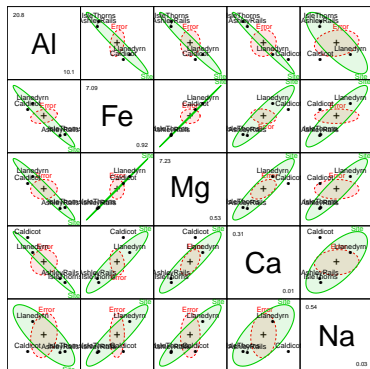
3 Reduced-rank displays

- Low-D displays of high-D data
- Canonical discriminant HE plots

4 Recent extensions

- Robust MLMs
- Influence diagnostics for MLMs
- Ridge regression plots

5 Conclusions



Slides: <http://datavis.ca/papers/user2019-2x2.pdf>

Overview: Research topics

Graphical methods for univariate response models well-developed. What about MLMs?

- This talk outlines research on graphical methods for **multivariate** linear models (MLMs)— extending visualization for multiple regression, ANOVA, and ANCOVA designs to those with several response variables.
- The topics addressed include:
 - Visualizing multivariate tests with **Hypothesis–Error (HE) plots** in 2D and 3D
 - Low-D views: Generalized canonical discriminant analysis → canonical HE plots
 - Visualization methods for tests of equality of covariance matrices in MANOVA designs
 - Extending these methods to **robust** MLMs
 - Developing multivariate analogs of **influence measures** and **diagnostic plots** for MLMs.

Overview: R packages

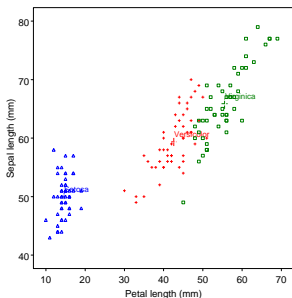
The following R packages implement these methods:

- **car** package: provides the infrastructure for hypothesis tests (`Anova()`) and tests of linear hypotheses (`linearHypothesis()`) in MLMs, including repeated measures designs.
- **heplots** package: implements the HE plot framework in 2D (`heplot()`), 3D (`heplot3d()`), and scatterplot matrix form (`pairs.mlm()`). Also provides:
 - `covEllipses()` for covariance ellipses, with optional robust estimation
 - `boxM()` and related methods for testing / visualizing equality of covariance matrices in MANOVA
 - Tutorial vignettes and many data set examples of use
- **candisc** package: generalized canonical discriminant analysis for an MLM, and associated plot methods.
- **mvinfluence** package: Multivariate extensions of leverage and influence (Cook's D) and `influencePlot.mlm()` in various forms.
- **genridge** package: Generalized 2D & 3D ridge regression plots.

Visual overview: Multivariate data, $Y_{n \times p}$

What we know how to do well (almost)

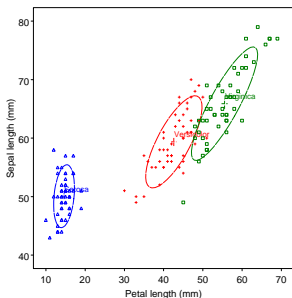
- 2 vars: Scatterplot



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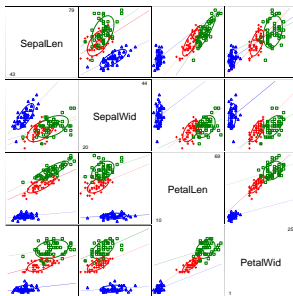
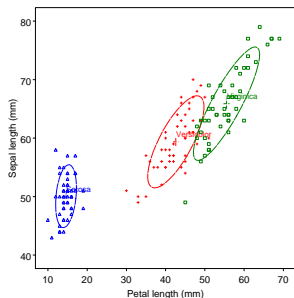
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- p vars: Scatterplot matrix (all pairs)



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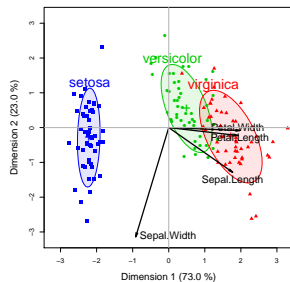
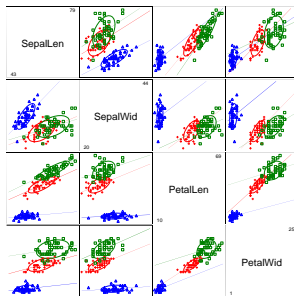
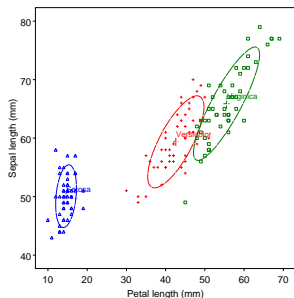
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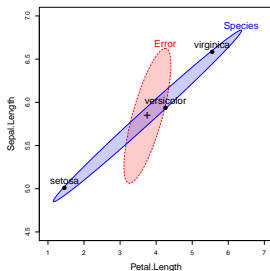


Visual overview: Multivariate linear model,

$$Y = XB + U$$

What is new here?

- 2 vars: HE plot— data ellipses of H (fitted) and E (residual) SSP matrices
- p vars: HE plot matrix (all pairs)

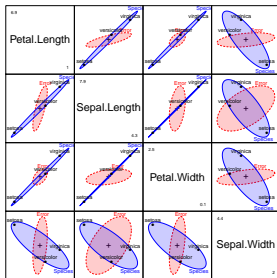
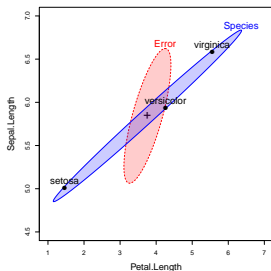


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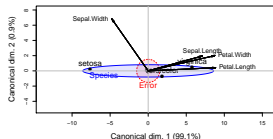
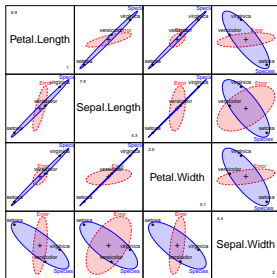
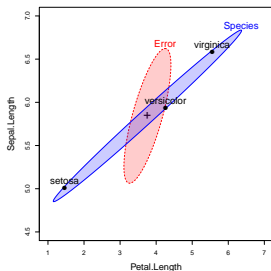


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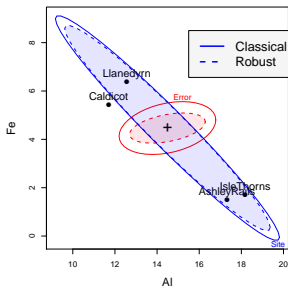
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Visual overview: Recent extensions

Extending univariate methods to MLMs:

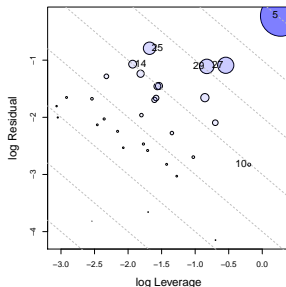
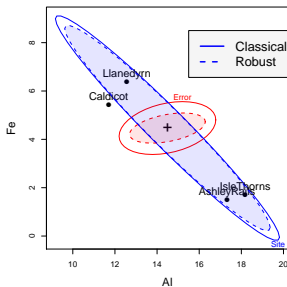
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- Influence measures and diagnostic plots for MLMs ([mvinfluence](#))



Visual overview: Recent extensions

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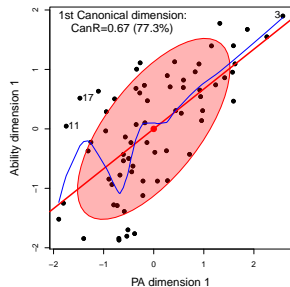
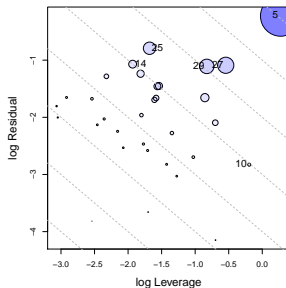
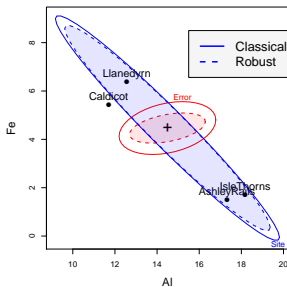
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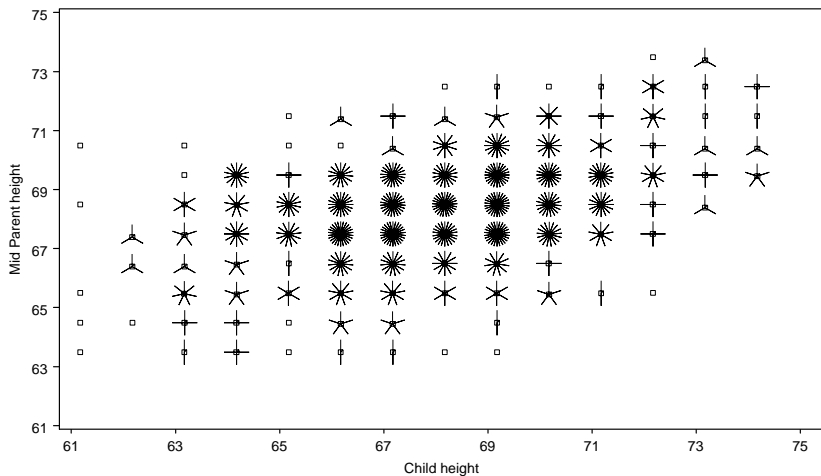
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Data ellipsoids: Visually sufficient summaries

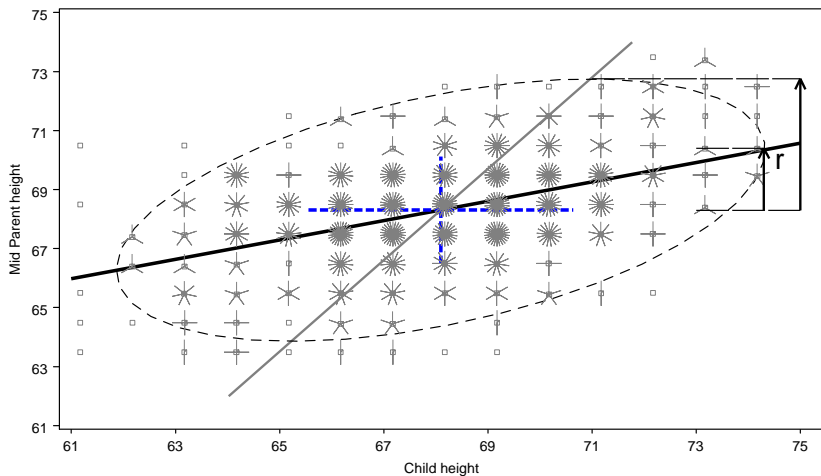
- For any p -variable, multivariate normal $\mathbf{y} \sim \mathcal{N}_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, the mean vector $\bar{\mathbf{y}}$ and sample covariance \mathbf{S} are **sufficient statistics**
- Geometrically, contours of constant density are **ellipsoids** centered at $\boldsymbol{\mu}$ with size and shape determined by $\boldsymbol{\Sigma}$
- \mapsto the **data** (concentration) ellipsoid, $\mathcal{E}(\bar{\mathbf{y}}, \mathbf{S})$ is a **sufficient visual summary**
- Easily robustified by using robust estimators of location and scatter

Data Ellipses: Galton's data



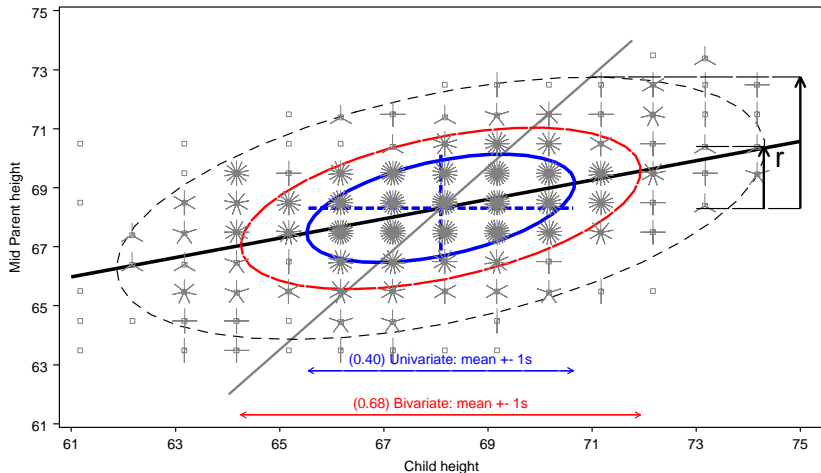
Galton's data on Parent & Child height

Data Ellipses: Galton's data



Data ellipse: Shows means, std. devs, regression lines, correlation

Data Ellipses: Galton's data



Radii: $c^2 = \chi_p^2(1 - \alpha)$ — 40%, 68% and 95% data ellipses

The Data Ellipse: Details

• Visual summary for bivariate relations

- **Shows:** means, standard deviations, correlation, regression line(s)
- **Defined:** set of points whose squared Mahalanobis distance $\leq c^2$,

$$D^2(\mathbf{y}) \equiv (\mathbf{y} - \bar{\mathbf{y}})^T \mathbf{S}^{-1} (\mathbf{y} - \bar{\mathbf{y}}) \leq c^2$$

\mathbf{S} = sample covariance matrix

- **Radius:** when \mathbf{y} is \approx bivariate normal, $D^2(\mathbf{y})$ has a large-sample χ_2^2 distribution with 2 degrees of freedom.
 - $c^2 = \chi_2^2(0.40) \approx 1$: 1 std. dev univariate ellipse– 1D shadows: $\bar{y} \pm 1s$
 - $c^2 = \chi_2^2(0.68) = 2.28$: 1 std. dev bivariate ellipse
 - $c^2 = \chi_2^2(0.95) \approx 6$: 95% data ellipse, 1D shadows: Scheffé intervals
- **Construction:** Transform the unit circle, $\mathcal{U} = (\sin \theta, \cos \theta)$,

$$\mathcal{E}_c = \bar{\mathbf{y}} + c\mathbf{S}^{1/2}\mathcal{U}$$

$\mathbf{S}^{1/2}$ = any “square root” of \mathbf{S} (e.g., Cholesky)

- **p variables:** Extends naturally to p -dimensional ellipsoids

The univariate linear model

- **Model:** $\mathbf{y}_{n \times 1} = \mathbf{X}_{n \times q} \boldsymbol{\beta}_{q \times 1} + \boldsymbol{\epsilon}_{n \times 1}$, with $\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I}_n)$
- **LS estimates:** $\hat{\boldsymbol{\beta}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$
- **General Linear Test:** $H_0 : \mathbf{C}_{h \times q} \boldsymbol{\beta}_{q \times 1} = \mathbf{0}$, where \mathbf{C} = matrix of constants; rows specify h linear combinations or contrasts of parameters.
- e.g., Test of $H_0 : \beta_1 = \beta_2 = 0$ in model $y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \epsilon_i$

$$\mathbf{C}\boldsymbol{\beta} = \begin{bmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$$

- All \rightarrow F-test: How big is SS_H relative to SS_E ?

$$F = \frac{SS_H / df_h}{SS_E / df_e} = \frac{MS_H}{MS_E} \rightarrow (MS_H - F MS_E) = 0$$

The multivariate linear model

- **Model:** $\mathbf{Y}_{n \times p} = \mathbf{X}_{n \times q} \mathbf{B}_{q \times p} + \mathbf{U}$, for p responses, $\mathbf{Y} = (\mathbf{y}_1, \mathbf{y}_2, \dots, \mathbf{y}_p)$
- **General Linear Test:** $H_0 : \mathbf{C}_{h \times q} \mathbf{B}_{q \times p} = \mathbf{0}_{h \times p}$
- Analogs of sums of squares, SS_H and SS_E are $(p \times p)$ matrices, \mathbf{H} and \mathbf{E}

$$\mathbf{H} = (\mathbf{C}\hat{\mathbf{B}})^T [\mathbf{C}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{C}^T]^{-1} (\mathbf{C}\hat{\mathbf{B}}) ,$$

$$\mathbf{E} = \mathbf{U}^T \mathbf{U} = \mathbf{Y}^T [\mathbf{I} - \mathbf{H}] \mathbf{Y} .$$

- Analog of univariate F is

$$\det(\mathbf{H} - \lambda \mathbf{E}) = 0 ,$$

- How big is \mathbf{H} relative to \mathbf{E} ?
 - Latent roots $\lambda_1, \lambda_2, \dots, \lambda_s$ measure the “size” of \mathbf{H} relative to \mathbf{E} in $s = \min(p, df_h)$ orthogonal directions.
 - Test statistics (Wilks' Λ , Pillai trace criterion, Hotelling-Lawley trace criterion, Roy's maximum root) all combine info across these dimensions

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- Visualizing H and E variation
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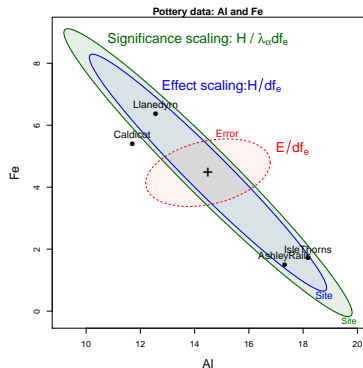
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Motivating Example: Romano-British Pottery

Tubb, Parker & Nicholson analyzed the chemical composition of 26 samples of Romano-British pottery found at four kiln sites in Britain.

- **Sites:** Ashley Rails, Caldicot, Isle of Thorns, Llanedryn
- **Variables:** aluminum (Al), iron (Fe), magnesium (Mg), calcium (Ca) and sodium (Na)
- → One-way MANOVA design, 4 groups, 5 responses

```
R> library(heplots)
```

```
R> Pottery
```

	Site	Al	Fe	Mg	Ca	Na
1	Llanedryn	14.4	7.00	4.30	0.15	0.51
2	Llanedryn	13.8	7.08	3.43	0.12	0.17
3	Llanedryn	14.6	7.09	3.88	0.13	0.20
	.	.	.			
25	AshleyRails	14.8	2.74	0.67	0.03	0.05
26	AshleyRails	19.1	1.64	0.60	0.10	0.03

Motivating Example: Romano-British Pottery

Questions:

- **Can** the content of Al, Fe, Mg, Ca and Na differentiate the sites?
- **How to understand** the contributions of chemical elements to discrimination?

Numerical answers:

```
R> pottery.mod <- lm(cbind(Al, Fe, Mg, Ca, Na) ~ Site)
R> car::Manova(pottery.mod)
```

```
Type II MANOVA Tests: Pillai test statistic
```

	Df	test	stat	approx	F	num	Df	den	Df	Pr(>F)
Site	3		1.55		4.30		15		60	2.4e-05 ***

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

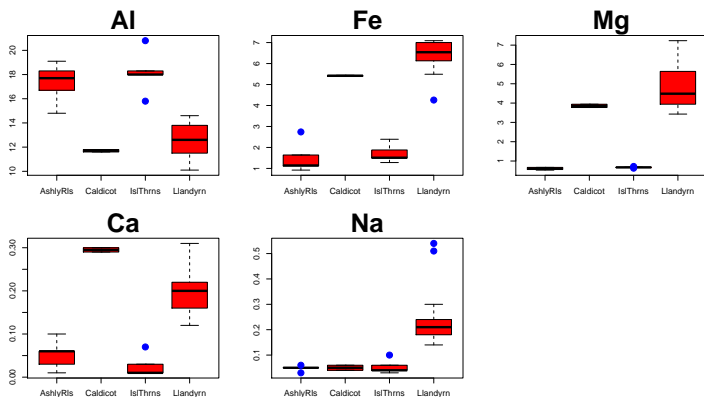
What have we learned?

- **Can**: YES! We can discriminate sites.
- But: **How to understand** the pattern(s) of group differences: ???

Motivating Example: Romano-British Pottery

Univariate plots are limited

- Do not show the *relations* of response variables to each other
- Do not show *how* variables contribute to multivariate tests

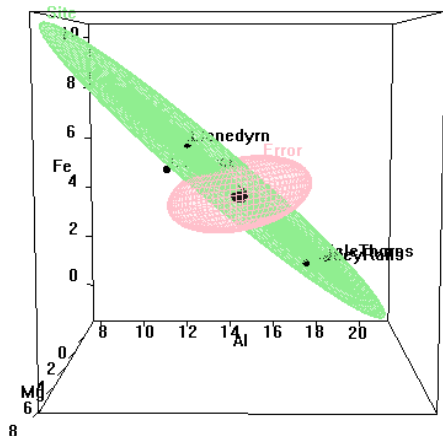


Motivating Example: Romano-British Pottery

Visual answer: HE plot

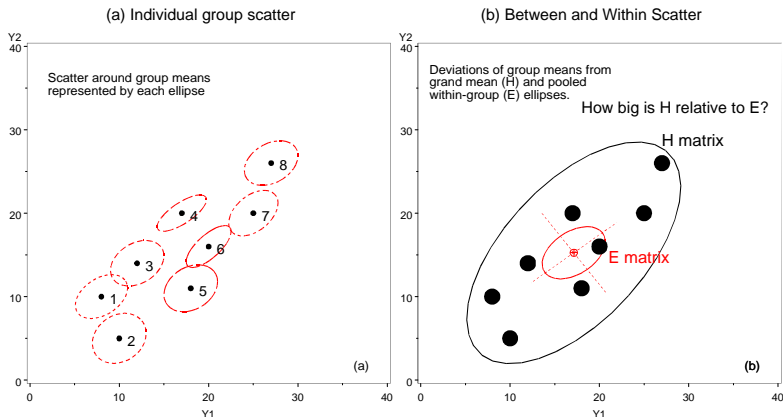
- Shows variation of means (H) relative to residual (E) variation
- Size and orientation of H wrt E : *how much* and *how* variables contribute to discrimination
- Evidence scaling: H is scaled so that it projects outside E iff null hypothesis is rejected.

Run heplot-movie.ppt



```
1 R> heplot3d(pottery.mod)
```

HE plots: Visualizing H and E variation



Ideas behind multivariate tests: (a) Data ellipses; (b) H and E matrices

- H ellipse: data ellipse for **fitted values**, $\hat{y}_{ij} = \bar{y}_j$.
- E ellipse: data ellipse of **residuals**, $\hat{y}_{ij} - \bar{y}_j$.

HE plot details: H and E matrices

Recall the data on 5 chemical elements in samples of Romano-British pottery from 4 kiln sites:

```
R> summary(Manova(pottery.mod))
```

```
Sum of squares and products for error:
```

	Al	Fe	Mg	Ca	Na
Al	48.29	7.080	0.608	0.106	0.589
Fe	7.08	10.951	0.527	-0.155	0.067
Mg	0.61	0.527	15.430	0.435	0.028
Ca	0.11	-0.155	0.435	0.051	0.010
Na	0.59	0.067	0.028	0.010	0.199

```
-----
```

```
Term: Site
```

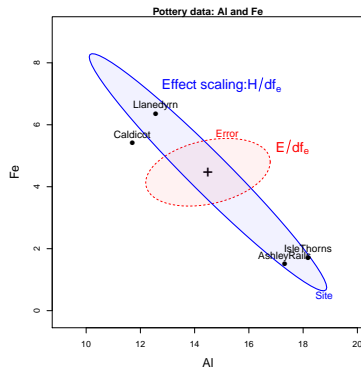
```
Sum of squares and products for hypothesis:
```

	Al	Fe	Mg	Ca	Na
Al	175.6	-149.3	-130.8	-5.89	-5.37
Fe	-149.3	134.2	117.7	4.82	5.33
Mg	-130.8	117.7	103.4	4.21	4.71
Ca	-5.9	4.8	4.2	0.20	0.15
Na	-5.4	5.3	4.7	0.15	0.26

- E matrix: Within-group (co)variation of residuals
 - diag: SSE for each variable
 - off-diag: \sim partial correlations
- H matrix: Between-group (co)variation of means
 - diag: SSH for each variable
 - off-diag: \sim correlations of means
- How big is H relative to E ?
- Ellipsoids: $\dim(H) = \text{rank}(H) = \min(p, df_h)$

HE plot details: Scaling H and E

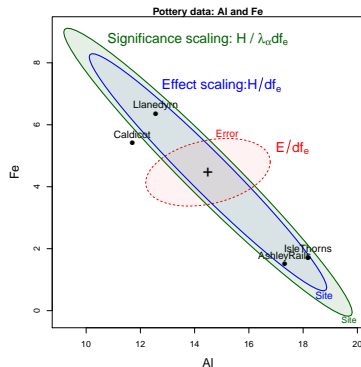
- The E ellipse is divided by $df_e = (n - p) \rightarrow$ data ellipse of residuals
 - Centered at grand means \rightarrow show factor means in same plot.
- “Effect size” scaling– $H/df_e \rightarrow$ data ellipse of fitted values.
- “Significance” scaling– H ellipse protrudes beyond E ellipse *iff* H_0 can be rejected by Roy maximum root test
 - $H/(\lambda_\alpha df_e)$ where λ_α is critical value of Roy’s statistic at level α .
 - direction of H wrt $E \mapsto$ linear combinations that depart from H_0 .



```
R> heplot(pottery.mod, size="effect")
```

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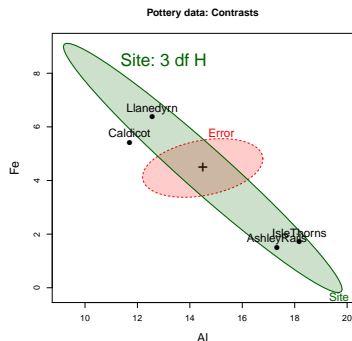


R> heplot(pottery.mod, size="evidence")

HE plot details: Contrasts and linear hypotheses

- An overall effect \mapsto an **H** ellipsoid of $s = \min(p, df_h)$ dimensions
- Linear hypotheses, of rank h ,
 $H_0 : \mathbf{C}_{h \times q} \mathbf{B}_{q \times p} = \mathbf{0}_{h \times p} \mapsto$ sub-ellipsoid of dimension h

$$\mathbf{C} = \begin{bmatrix} 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \end{bmatrix}$$

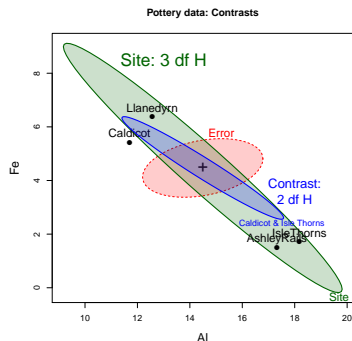


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- 1D tests and contrasts \mapsto degenerate 1D ellipses (lines)
- Beautiful geometry:
 - Sub-hypotheses are **tangent** to enclosing hypotheses
 - Orthogonal contrasts form **conjugate axes**

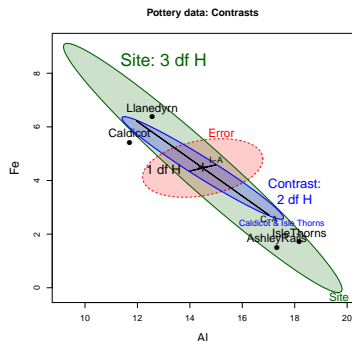


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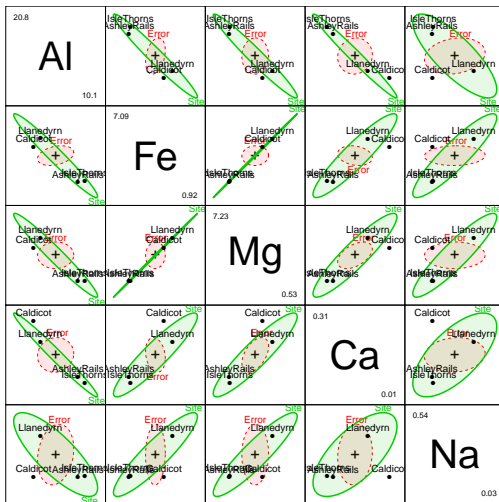
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HE plot matrices: All bivariate views

AL stands out –
opposite pattern
 $r(\overline{Fe}, \overline{Mg}) \approx 1$

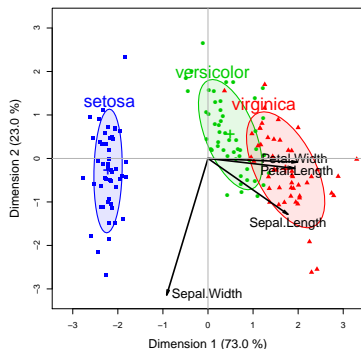
▶ Jump to low-D



R> pairs(pottery.mod)

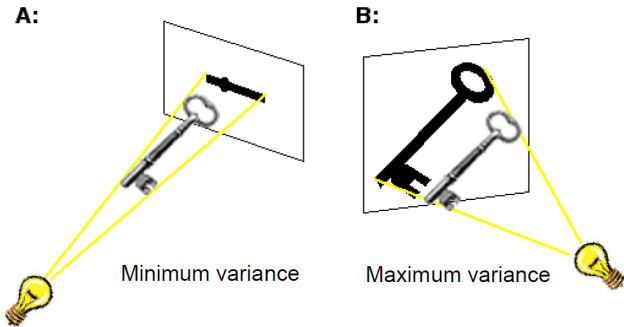
Outline

- 1 Background
 - Overview
 - Visual overview
 - Data ellipses
 - The Multivariate Linear Model
- 2 Hypothesis Error (HE) plots
 - Motivating example
 - Visualizing H and E variation
 - MANOVA designs
- 3 **Reduced-rank displays**
 - Low-D displays of high-D data
 - Canonical discriminant HE plots
- 4 Recent extensions
 - Robust MLMs
 - Influence diagnostics for MLMs
 - Ridge regression plots
- 5 Conclusions



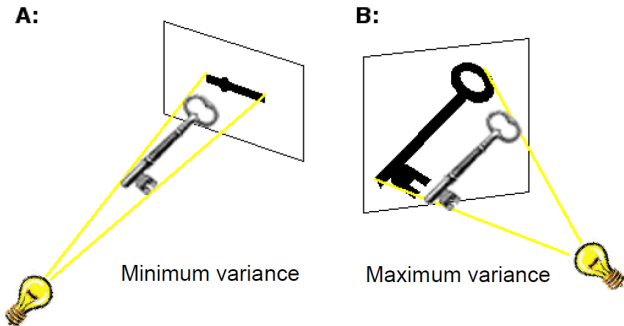
Low-D displays of high-D data

- High-D data often shown in 2D (or 3D) views— orthogonal projections in variable space— **scatterplot**
- **Dimension-reduction** techniques: project the data into subspace that has the largest *shadow*— e.g., accounts for largest variance.
- → low-D approximation to high-D data



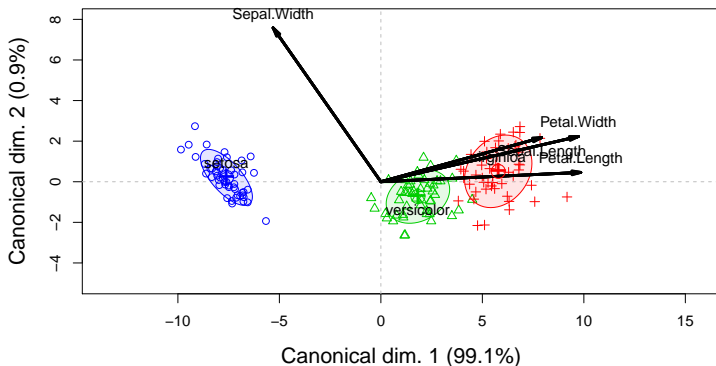
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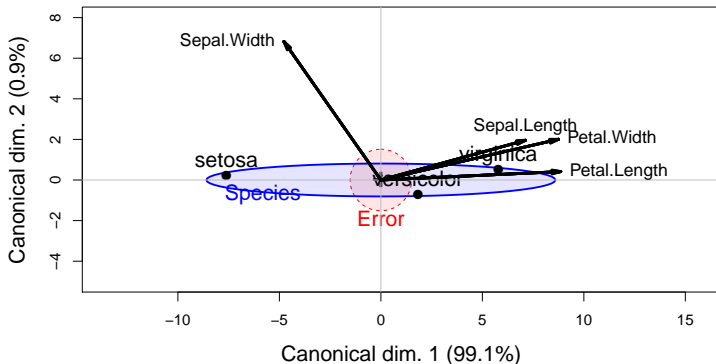
Canonical discriminant HE plots

- As with biplot, we can visualize MLM hypothesis variation for *all* responses by projecting \mathbf{H} and \mathbf{E} into low-rank space.
- Canonical projection:** $\mathbf{Y}_{n \times p} \mapsto \mathbf{Z}_{n \times s} = \mathbf{Y}\mathbf{E}^{-1/2}\mathbf{V}$, where \mathbf{V} = eigenvectors of $\mathbf{H}\mathbf{E}^{-1}$.
- This is the view that maximally discriminates among groups, ie max. \mathbf{H} wrt \mathbf{E} !



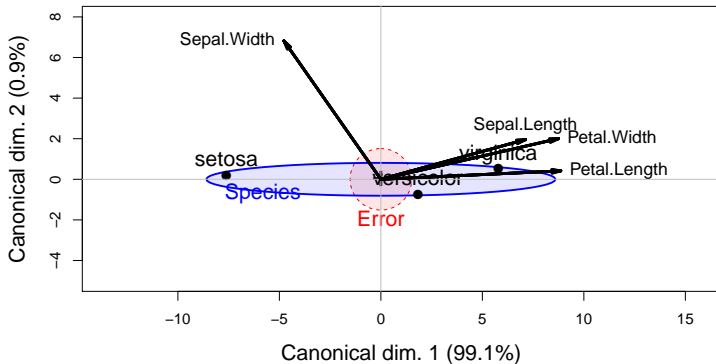
Canonical discriminant HE plots

- Canonical HE plot is just the HE plot of canonical scores, $(\mathbf{z}_1, \mathbf{z}_2)$ in 2D,
- or, $\mathbf{z}_1, \mathbf{z}_2, \mathbf{z}_3$, in 3D.
- As in biplot, we add vectors to show relations of the \mathbf{y}_i response variables to the canonical variates.
- variable vectors here are **structure coefficients** = correlations of variables with canonical scores.



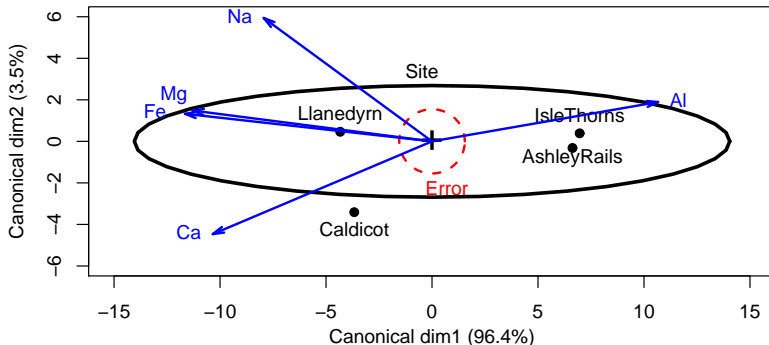
Canonical discriminant HE plots: Properties

- Canonical variates are uncorrelated: \mathbf{E} ellipse is spherical
- \mapsto axes must be equated to preserve geometry
- Variable vectors show how variables discriminate among groups
- Lengths of variable vectors \sim contribution to discrimination



Canonical discriminant HE plots: Pottery data

- Canonical HE plots provide 2D (3D) visual summary of H vs. E variation
- Pottery data: $p = 5$ variables, 4 groups $\mapsto df_H = 3$
- Variable vectors: Fe, Mg and Al contribute to distinguishing (Caldicot, Llandryn) from (Isle Thorns, Ashley Rails): 96.4% of mean variation
- Na and Ca contribute an additional 3.5%. **End of story!**



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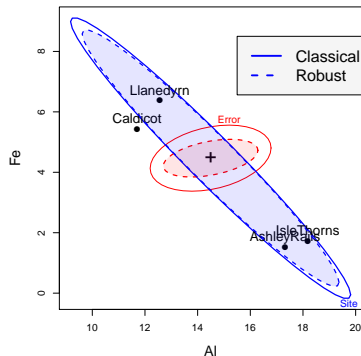
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Robust MLMs

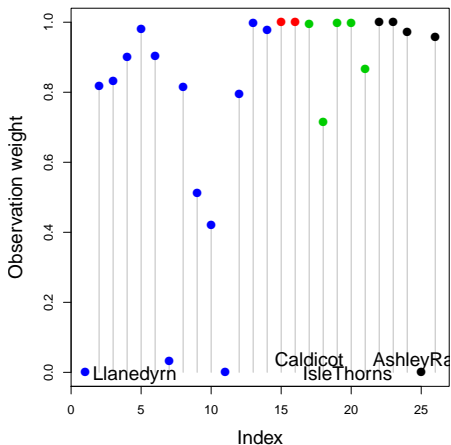
- R has a large collection of packages dealing with robust estimation:
 - `robust::lmrob()`, `MASS::rlm()`, for *univariate* LMs
 - `robust::glmrob()` for univariate *generalized* LMs
 - **High breakdown-bound** methods for robust *PCA* and robust covariance estimation
 - However, none of these handle the **fully general MLM**
- `heplots` now provides `robmlm()` for robust MLMs:
 - Uses a simple M-estimator via iteratively re-weighted LS.
 - Weights: calculated from Mahalanobis squared distances, using a simple robust covariance estimator, `MASS::cov.trob()` and a weight function, $\psi(D^2)$.
$$D^2 = (\mathbf{Y} - \hat{\mathbf{Y}})^T \mathbf{S}_{\text{trob}}^{-1} (\mathbf{Y} - \hat{\mathbf{Y}}) \sim \chi_p^2 \quad (1)$$
 - This fully extends the `"mlm"` class
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Robust MLMs: Example

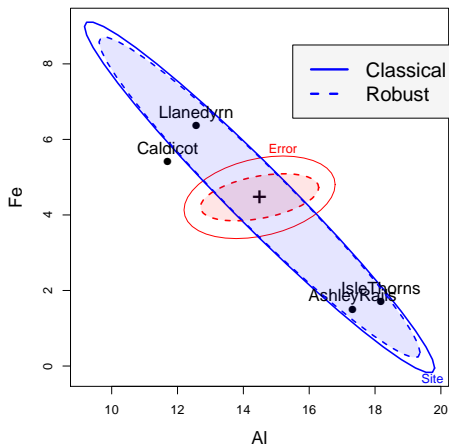
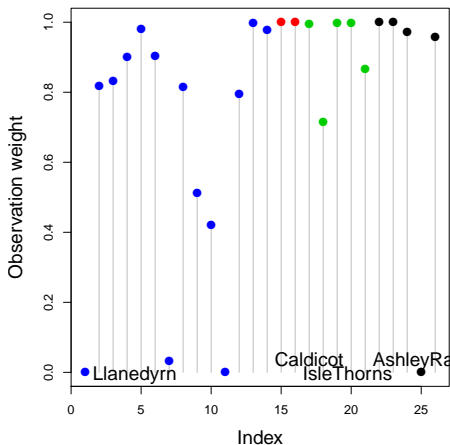
For the Pottery data:



- Some observations are given weights ~ 0
- The E ellipse is considerably reduced, enhancing apparent significance

Robust MLMs: Example

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Influence diagnostics for MLMs

- Influence measures & diagnostic plots well-developed for *univariate* LMs
 - Influence measures: Cook's D, DFFITS, dfbetas, etc.
 - Diagnostic plots: Index plots, `car::influencePlot()` for LMs
 - However, these have been unavailable for MLMs
- The `mvinfluence` package now provides:
 - Calculation for multivariate analogs of univariate influence measures (following Barrett & Ling, 1992), e.g., Hat values & Cook's D :

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$$D_I = [\text{vec}(\mathbf{B} - \mathbf{B}_{(I)})]^T [\mathbf{S}^{-1} \otimes (\mathbf{X}^T \mathbf{X})][\text{vec}(\mathbf{B} - \mathbf{B}_{(I)})] \quad (3)$$

- Provides deletion diagnostics for *subsets* (I) of size $m \geq 1$.
- e.g., $m = 2$ can reveal cases of *masking* or *joint influence*.
- Extension of `influencePlot()` to the multivariate case.
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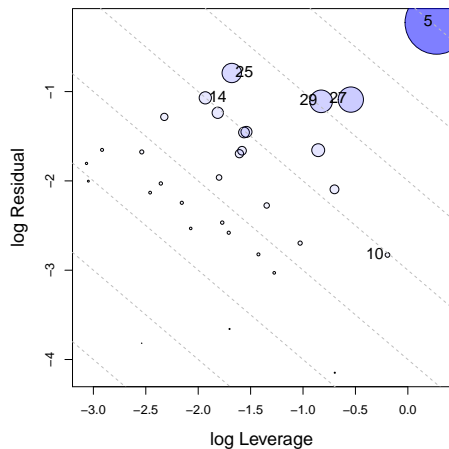
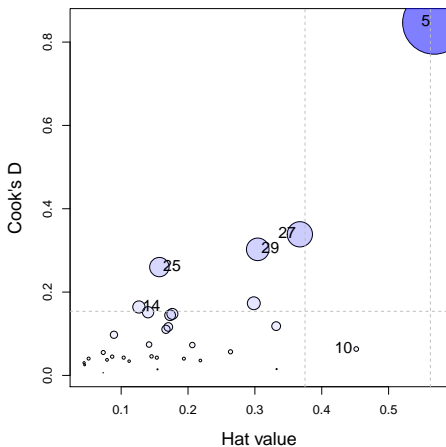
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For the Rohwer data:

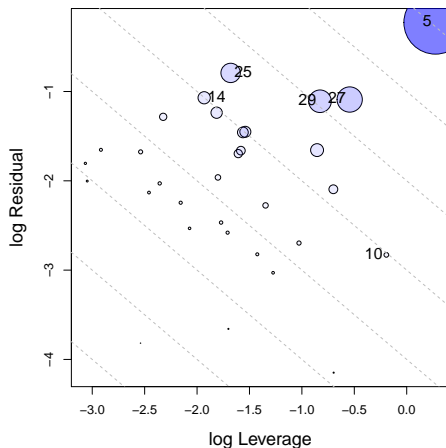


Cook's D vs. generalized Hat value

Leverage - Residual (LR) plot

Influence diagnostics for MLMs: LR plots

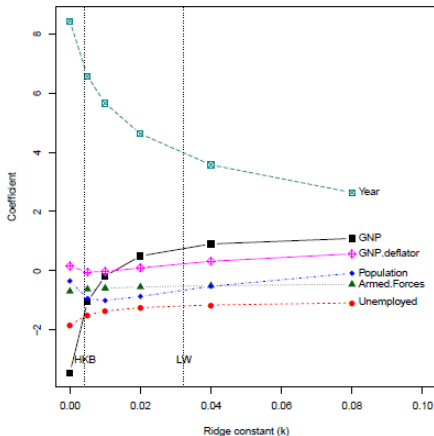
- Main idea: Influence \sim Leverage (L) \times Residual (R)
- $\mapsto \log(\text{Infl}) = \log(L) + \log(R)$
- \mapsto contours of constant influence lie on lines with slope = -1.
- Bubble size \sim influence (Cook's D)
- This simplifies interpretation of influence measures



Ridge regression plots

Shrinkage methods often use **ridge trace plots** to visualize effects

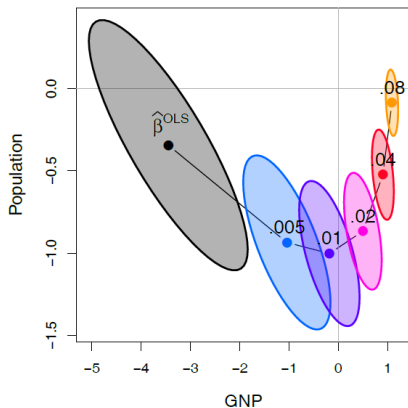
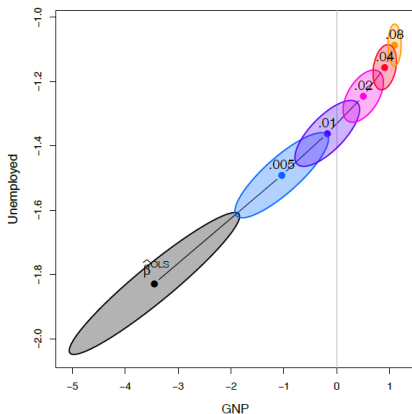
- Typical: univariate line plot of β_k vs. shrinkage, k
- What can you see here regarding bias vs. precision?
- This is the **wrong graphic form**, for a **multivariate** problem!
- Goal: visualize $\widehat{\beta}_k$ vs. $\widehat{\text{Var}}(\widehat{\beta}_k)$



Generalized ridge trace plots

Rather than plotting just the univariate trajectories of β_k vs. K , plot the 2D (3D) confidence **ellipsoids** over the same range of k .

- Centers of the ellipsoids are $\widehat{\beta}_k$ – same info as in univariate plot.
- Can see how change in one coefficient is related to changes in others.
- Relative size & shape of ellipsoids show **directly** effect on precision.



Conclusions: Graphical methods for MLMs

Summary & Opportunities

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 - Useful for multiple-group, MANOVA data
 - Embed in scatterplot matrix: pairwise, bivariate relations
 - Easily extend to show partial relations, robust estimators, etc.
- **HE plots**: visual summary of multivariate tests for MANOVA and MMRA
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— FIN et Merci —

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